

# Econometrics

## Chapter 14: Panel Data Methods

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# Fixed Effects Estimation

- When there are unobserved fixed effects, an alternative to first differences is fixed effects estimation.
- Consider the average over time of

$$y_{it} = \beta_0 + \beta_1 x_{it1} + \dots + \beta_k x_{itk} + \mu_i + u_{it}.$$

The average of  $\mu_i$  will be  $\bar{\mu}_i$ , so if you subtract the mean,  $\mu_i$  will be differenced out just as when doing first differences. That is,

$$y_{it} - \bar{y}_i = \beta_1 (x_{it1} - \bar{x}_{i.1}) + \dots + \beta_k (x_{itk} - \bar{x}_{i.k}) + u_{it} - \bar{u}_i. \quad (1)$$

- This method is also identical to including a separate intercept for every individual.

# First Differences vs Fixed Effects

- First Differences and Fixed Effects will be exactly the same when  $T = 2$ . Why?
- For  $T > 2$ , the two methods are different.
- Fixed effects estimation is easily implemented for unbalanced panels ( $T$  different for each individual), not just for balanced panels.

- Start with the same basic model with a composite error

$$y_{it} = \beta_0 + \beta_1 x_{it1} + \dots + \beta_k x_{itk} + \mu_i + u_{it}.$$

- Previously we assumed that  $\mu_i$  was correlated with the  $x$ 's, but what if it's not?
- OLS would be consistent in that case, but composite errors will be serially correlated.
- Need to transform the model and do GLS to solve the problem and make correct inferences.
- The GLS estimator is called the random effects estimator.

- End up with a sort of weighted average of OLS and Fixed Effects – use quasi-demeaned data.
- Let

$$\lambda = 1 - \sqrt{\frac{\sigma_u^2}{\sigma_u^2 + T\sigma_a^2}}$$

- The RE estimator is obtained by running regression on

$$y_{it} - \lambda \bar{y}_i = \beta_1(x_{it1} - \lambda \bar{x}_{i.1}) + \dots + \beta_k(x_{itk} - \lambda \bar{x}_{i.k}) + v_{it} - \lambda \bar{v}_i,$$

where  $v_{it} = \mu_i + u_{it}$ .

- If  $\lambda = 1$ , then this is just the fixed effects estimator.
- So, the bigger the variance of the unobserved effect, the closer it is to FE.

## Example

Three Different Estimators of a Wage Equation (WAGEPAN.RAW;  
Wooldridge p.472)

## Example

Dependent Variable:  $\log(wage)$

Independent Variables	Pooled OLS	Random Effects	Fixed Effects
<i>educ</i>	.091 (.005)	.092 (.011)	
<i>black</i>	-.139 (.024)	-.139 (.048)	
<i>hispan</i>	.016 (.021)	.022 (.043)	
<i>exper</i>	.067 (.014)	.106 (.015)	
<i>exper</i> <sup>2</sup>	-.0024 (.0008)	-.0047 (.0007)	-.0052 (.0007)
<i>married</i>	.108 (.016)	.064 (.017)	.047 (.018)
<i>union</i>	.182 (.017)	.106 (.018)	.080 (.019)

## Example

(continued) Marriage premium falls to 4.7% when fixed effects estimator is used. This is consistent with the idea that marriage and unobserved ability are positively correlated. After taking into account individual ability, marriage still carries a premium. Why? Married people may be more stable in work places and work harder. Employers may be willing to pay for a premium for these possible reasons.



# Fixed Effects or Random?

- If the individual effects and regressors are correlated, the FE estimator should be used. If not, the RE estimator is BLUE.
- Hausman's (1978) test can be used to find out whether the individual effects and regressors are correlated or not.
- The test statistic is based on FE-estimator - RE-estimator for the null hypothesis that the individual effects and regressors are not correlated.

# Other Uses of Panel Methods

- It is possible to think of models where there is an unobserved fixed effect, even if we do not have true panel data.
- A well-known example is an unobserved family effect.
- See  
Ashenfelter and Krueger (1994), Estimates of the Economic Return to Schooling from a New Sample of Twins, *American Economic Review*, pp. 1157-1173.
- Can difference siblings.
- Can estimate family fixed effect model.

## Other Uses of Panel Methods

Let

$y_{1i}, y_{2i}$  : Logarithms of the wage rates of the first and second twin in the  $i$ -th couple

$X_i$  : The set of variables that vary by family, but not across twins (age, race, measures of family background)

$Z_{1i}, Z_{2i}$  : The set of variables that may vary across the twins (education levels, union status, job tenure, marital status)

$\mu_i$  : An unobservable components that vary by family.

- The model is

$$\begin{aligned}y_{1i} &= \alpha' X_i + \beta' Z_{1i} + \mu_i + \varepsilon_{1i} \\y_{2i} &= \alpha' X_i + \beta' Z_{2i} + \mu_i + \varepsilon_{2i}.\end{aligned}\tag{2}$$

Here  $\beta$  denotes structural effects of the observables on earnings.

- Differencing the two equations in (2) gives

$$y_{1i} - y_{2i} = \beta'(Z_{1i} - Z_{2i}) + \varepsilon_{1i} - \varepsilon_{2i}.$$

We may run OLS or IV on this equation. The IV estimation is considered due to the possible measurement errors in  $Z_{1i}$  and  $Z_{2i}$ .

- Many of the theories we already know about both cross section and time series data can be applied to panel data.
- Can test and correct for serial correlation in the errors.
- Can test and correct for heteroskedasticity.
- Can estimate standard errors robust to both.