

# Econometrics

## Chapter 9: Heteroskedasticity

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# What is Heteroskedasticity?

- The assumption of homoskedasticity means that  $E(u_t^2 \mid \text{all } x's) = \sigma^2$  ( $\sigma^2$  is a constant that does not change over  $i$ ).
- If this is not true, that is if the variance of  $u_t$  is different for different values of the  $x$ 's, then the errors are heteroskedastic.

# What is Heteroskedasticity?

## Example

(Cross-section consumption function)

$y_i$  : household  $i$ 's consumption

$x_i$  : household  $i$ 's income

Consider the regression equation

$$y_i = \beta_0 + \beta_1 x_i + u_i.$$

Then,  $\text{Var}(y_i | x_i) = \text{Var}(u_i | x_i)$ . It is highly likely that  $\text{Var}(y_i | x_i)$  is large when  $x_i$  is large.

# Why Worry About Heteroskedasticity?

- OLS is still unbiased and consistent, even if we do not assume homoskedasticity.

## Example

Consider the model

$$y_t = \beta_0 + \beta_1 x_t + u_t, \quad E(u_t) = 0 \text{ and } E(u_t^2) = \sigma_t^2.$$

Assume for simplicity that  $\{x_t\}$  is a sequence of constants. Recall that

$$\hat{\beta}_1 - \beta_1 = \sum_{t=1}^n w_t u_t,$$

where  $w_t = \frac{x_t - \bar{x}}{\sum_{t=1}^n (x_t - \bar{x})^2}$ . Then,  $E(\hat{\beta}_1) = \beta_1$  since

$$E(\hat{\beta}_1 - \beta_1) = E\left(\sum_{t=1}^n w_t u_t\right)$$

# Why Worry About Heteroskedasticity?

- Furthermore, if we assume  $\sigma_t^2 < M$  for all  $t$ ,

$$\begin{aligned} \text{Var}(\hat{\beta}_1) &= E \left[ (\hat{\beta}_1 - \beta_1)^2 \right] \\ &= E \left[ \left( \sum_{t=1}^n w_t u_t \right)^2 \right] \\ &= \sum_{t=1}^n w_t^2 (E u_t^2) \\ &= \sum_{t=1}^n w_t^2 \sigma_t^2 \\ &= \frac{\sum_{t=1}^n (x_t - \bar{x})^2 \sigma_t^2}{\left( \sum_{t=1}^n (x_t - \bar{x})^2 \right)^2} \leq \frac{M}{\sum_{t=1}^n (x_t - \bar{x})^2}. \end{aligned}$$

Thus, if  $\sum_{t=1}^n (x_t - \bar{x})^2 \rightarrow \infty$  as  $n$  increases, the estimator is consistent.

# Why Worry About Heteroskedasticity?

- The usual formula for standard errors should be different if we have heteroskedasticity. Thus, we can not use the usual t or F test for drawing inferences.

- Consider the simple linear regression model. Then,  $\text{Var}(\hat{\beta}_1)$  can be estimated by

$$\frac{\sum_{t=1}^n (x_t - \bar{x})^2 \hat{u}_t^2}{(\sum_{t=1}^n (x_t - \bar{x})^2)^2}.$$

- Adjusted t-ratio is defined by

$$\frac{\hat{\beta}_1}{\sqrt{\frac{\sum_{t=1}^n (x_t - \bar{x})^2 \hat{u}_t^2}{(\sum_{t=1}^n (x_t - \bar{x})^2)^2}}}.$$

This has a standard normal distribution in the limit. If this is used, heteroskedasticity is taken care of.

- Write

$$\sqrt{n}(\hat{\beta}_1 - \beta_1) = \left( \frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 \right)^{-1} \frac{1}{\sqrt{n}} \sum_{t=1}^n (x_t - \bar{x}) u_t.$$

The variance of  $\frac{1}{\sqrt{n}} \sum_{t=1}^n (x_t - \bar{x}) u_t$  is  $\frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 \sigma_t^2$ .



- By the WLLN,

$$\frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 u_t^2 - \frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 \sigma_t^2 \xrightarrow{p} 0.$$

Since  $u_t^2 - \hat{u}_t^2 \xrightarrow{p} 0$ ,  $\frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 \sigma_t^2$  can be estimated by  $\frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 \hat{u}_t^2$ .

- Thus, if  $\lim \frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 = M$  and  $\lim \frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 \sigma_t^2 = L$ ,

$$\frac{\hat{\beta}_1 - \beta_1}{\sqrt{\frac{\sum_{t=1}^n (x_t - \bar{x})^2 \hat{u}_t^2}{(\sum_{t=1}^n (x_t - \bar{x})^2)^2}}} = \frac{\sqrt{n} (\hat{\beta}_1 - \beta_1)}{\sqrt{\frac{\frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2 \hat{u}_t^2}{(\frac{1}{n} \sum_{t=1}^n (x_t - \bar{x})^2)^2}}} \xrightarrow{d} \frac{N(0, M^{-2}L)}{\sqrt{\frac{L}{M^2}}} = N(0, 1)$$

- Important to remember that these robust standard errors only have asymptotic justification. In small samples, the t-statistic formed with the robust standard error may not have a distribution close a standard normal distribution, and inferences may not not be correct.

# Generalized least squares (GLS) estimation

- Model

$$y_t = \beta_0 + \beta_1 x_{1t} + \cdots + \beta_k x_{kt} + u_t.$$

If  $\text{Var}(u_t) = \sigma_t^2$ , consider the transformed model

$$y_t/\sigma_t = \beta_0/\sigma_t + \beta_1 x_{1t}/\sigma_t + \cdots + \beta_k x_{kt}/\sigma_t + u_t/\sigma_t.$$

- The error term  $u_t/\sigma_t$  satisfy classical assumptions. Thus, the OLS estimator from this transformed model is BLUE. This estimator is called the generalized least squares (GLS) estimator.
- In practice,  $\sigma_t$  is not known.