

Econometrics

Method-of-Moments and Maximum Likelihood Estimators

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Method of Moments Estimators

- Let X_1, \dots, X_n be an iid sample from a population with pdf $f(x | \theta_1, \dots, \theta_k)$.
- Let $\mu'_j = EX^j$ ($j = 1, \dots, k$). The population moments are typically a function of $(\theta_1, \dots, \theta_k)$.
- Define

$$m_j = \frac{1}{n} \sum_{i=1}^n X_i^j, \quad (j = 1, \dots, k).$$

The MM estimator is obtained by the following system of equations

$$m_j = \mu'_j(\theta_1, \dots, \theta_k), \quad (j = 1, \dots, k).$$

Example

Suppose $X_i \sim iid N(\theta, \sigma^2)$ ($i = 1, \dots, n$). We have $m_1 = \bar{X}$ and $m_2 = \frac{1}{n} \sum_{i=1}^n X_i^2$, $\mu'_1 = \theta$ and $\mu'_2 = \theta^2 + \sigma^2$. The MM estimators are obtained by solving the equations

$$\begin{aligned}\bar{X} &= \theta; \\ \frac{1}{n} \sum_{i=1}^n X_i^2 &= \theta^2 + \sigma^2.\end{aligned}$$

Thus $\hat{\theta} = \bar{X}$ and $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$.

Example

Suppose that $\{X_t, t = 1, \dots, n\}$ is a stationary process. Then,

$$\begin{aligned}EX_t &= m \\E(X_t - m)(X_{t+h} - m) &= EX_t X_{t+h} - m^2 = \gamma(h).\end{aligned}$$

The MM estimators of m and $\gamma(h)$ are obtained by solving the equations

$$\begin{aligned}\bar{X} &= m \\n^{-1} \sum_{t=1}^{n-h} (X_t - m)(X_{t+h} - m) &= \gamma(h).\end{aligned}$$

Maximum Likelihood Estimators

- Let X_1, \dots, X_n be an iid sample from a population with pdf $f(x | \theta_1, \dots, \theta_k)$. The likelihood function is defined by

$$\begin{aligned}L(\theta \mid \mathbf{x}) &= L(\theta_1, \dots, \theta_k \mid x_1, \dots, x_n) \\ &= \prod_{i=1}^n f(x_i \mid \theta_1, \dots, \theta_k).\end{aligned}$$

- For each sample point \mathbf{x} , let $\hat{\theta}(\mathbf{x})$ be a parameter value at which $L(\theta \mid \mathbf{x})$ attains its maximum as a function of θ , with \mathbf{x} held fixed. A maximum likelihood estimator (MLE) of the parameter θ based on a sample \mathbf{x} is $\hat{\theta}(\mathbf{x})$.

Maximum Likelihood Estimators

- The MLE is the parameter point for which the observed sample is most likely.
- The MLE is a good point estimator, possessing some of the optimality properties.

Example

Suppose $X_i \sim iid N(\theta, 1)$ ($i = 1, \dots, n$). Then,

$$L(\theta | \mathbf{x}) = \prod_{i=1}^n \frac{1}{(2\pi)^{1/2}} e^{-(x_i - \theta)^2 / 2} = \frac{1}{(2\pi)^{n/2}} e^{-\sum_{i=1}^n (x_i - \theta)^2 / 2}.$$

The first-order condition gives

$$\sum_{i=1}^n (x_i - \theta) = 0,$$

giving $\hat{\theta} = \bar{x}$. The second-order condition ($\frac{\partial^2}{\partial \theta^2} L(\theta | \mathbf{x})|_{\theta=\bar{x}} < 0$) is also satisfied. Thus, \bar{x} is the only extreme point in the interior and it yields a maximum. At the boundaries $\theta = \pm\infty$, $L(\theta | \mathbf{x}) = 0$. Thus, \bar{x} yields the global maximum likelihood.

Example

(Bernoulli MLE) $X_i \sim iid B(1, p)$ ($i = 1, \dots, n$). The likelihood function is

$$L(p \mid \mathbf{x}) = \prod_{i=1}^n p^{x_i} (1-p)^{1-x_i} = p^y (1-p)^{n-y},$$

where $y = \sum_{i=1}^n x_i$. The log likelihood function is

$$\log L(p \mid \mathbf{x}) = y \log p + (n-y) \log(1-p).$$

If $0 < y < n$, $\frac{\partial}{\partial p} L(p \mid \mathbf{x}) = 0$ gives $\hat{p} = y/n$. If $y = 0$ or $y = n$,

$$\log L(p \mid \mathbf{x}) = \begin{cases} n \log(1-p) & \text{if } y = 0 \\ n \log p & \text{if } y = n \end{cases}.$$

Because $n \log(1-p)$ and $n \log p$ are maximized at $p = 0$ and $p = 1$, $\hat{p} = y/n$ is the MLE in either case.